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- What are the risk factors affecting graft failure?
- > Can clinical parameters predicting late graft failure be identified?
- > Do the curve changes of clinical parameters indicating kidney function predict graft failure?

Introduction

Chronic kidney disease (CKD) is a condition in which the kidneys gradually lose their function due to damage, eventually leading to the need for dialysis or kidney transplantation (Tx). Although a kidney transplant is a treatment, it is not a permanent solution; after a certain period, the patient may need to undergo dialysis or another transplant. Globally, the prevalence of CKD is estimated at 9.1% (Cockwell & Fisher, 2020). In our country, its prevalence is 15.7%. It ranks as the twelfth leading cause of death worldwide (Cockwell & Fisher, 2020).

The transplanted kidney has a certain lifetime and it depends on recipient and donor related, and many other clinical factors. Considering graft failure risk, patients are followed over years to monitor their kidney function during the post-transplantation period. This longitudinal observational period provides longitudinal and survival outcomes which are related and allow to model graft failure mechanism.

In this study, the relationship between kidney survival and temporal changes in creatinine and hematocrit, which are kidney function biomarkers, was examined with multivariate joint models.



Method

Study design

Retrospective cohort

Study data

The application data used in this study was obtained by applying to the Non-Interventional Ethics Committee of Acıbadem Bursa Hospital. There is cross-sectional and longitudinal data (~88,000 rows) for a total of 1002 patients who underwent kidney transplantation between December 2011 and January 2021.

Study outcomes

Survival Outcome Variables:

Graft failure: Defined as the absence of kidney function occurring at any time after transplantation due to irreversible graft damage that requires chronic dialysis and/or re-transplantation, or the patient's death.

Longitudinal Outcome Variables:

Serum creatinine level (CRE: mg/dl) Hematocrit level (HCT: %) Proteinuria (Positive/Negative)

Independent Variables of the Study

•Gender, Age at transplantation, Donor type, Total number of mismatches, Dialysis type and duration, History of any type of rejection, Etiology types

Statistical Models

A significance level of 0.05 was used in all analyses. R software (R software, version 4.0.5, packages: arsenal, JMbayes2 for joint modeling, survival for proportional hazard tests, MuMIn for mixed model R² calculation, GLMMadaptive, R Foundation for Statistical Computing, Vienna, Austria; <u>http://r-project.org</u>) was used.

Survival Submodel: The variable selection for the Cox regression model was based on clinical knowledge and the completeness and accuracy of the data content. The Cox regression assumption was examined using the proportional hazard Schoenfeld residual test. Since the baseline hazard function in joint models must be defined parametrically, a B-spline was used within the joint model.

Longitudinal Submodel: The time-dependent variation of numerical longitudinal variables was modeled using a natural spline with four knots. In this model, fixed effects for age, gender, and donor type, along with random intercepts and random slopes, were included. The conditional variance explained by the model (fixed effects + random effects) and the normality of the residuals were checked. For proteinuria, a generalized linear mixed model (with a binomial logit link function) was used, including random intercepts and slopes with only the fixed effect of time.

Joint Model: Due to the difficulty in calculating the maximum likelihood estimator in multivariate joint models, estimation was performed using a Bayesian approach.

Model results and model selection were based on the Deviance Information Criterion (DIC), trace plots, and posterior density plots.

Flow chart





<u>Results</u>

Recipient characteristics

	(N=749)	
Gender		CKD etiology
Male	475 (63.4%)	Glomerulonephritis
Female	274 (36.6%)	Seconder Glomerulon
Age		Tubulaintenstitial namb
Mean (SD)	44.5 (11.9)	iubulointerstitiai nepr
Median (Min-Max)	46.0 (18.0 - 79.0)	Polycystic kidney
BIVII	10	Genetic
Moon (SD)		Diabetes
Median (Min-Max)	25.4 (4.7)	Neurogenic Bladder
Blood type	24.8 (11.8 - 47.0)	Vesicoureteral reflux
A	334 (44.6%)	Dialysis type
AB	64 (8.5%)	Dariton col dialucia /
В	139 (18.6%)	Peritoneal dialysis (P
0	212 (28.3%)	Hemodialysis (HD)
Second Tx	15 (2.0%)	HD+PD
Smoking		Preemptive
Yes	220 (29.4%)	Total dialysis duration
No	49 (6.5%)	Mean (SD)
Former-smoker	116 (15.5%)	Median (Min-May)
Unknown	364 (48.6%)	
Alcohol		Pre-IX CRE
Yes	358 (47.8%)	Missing
No Fama an daialaan	6 (0.8%)	Mean (SD)
Former-drinker	21 (2.8%)	Median (Min-Max)
Coronary artery disease	364 (48.6%)	Celluluar Rejection
(CAD)		Humoral Rejection
Missing	2	Hyperacute Rejection
Yes	238 (31.9%)	Dejection
No	147 (19.7%)	Rejection
Unknown	362 (48.5%)	Ex
Hypertension (HT)		Graft loss
Missing	2	Follow up duration (y
Yes	127 (17.0%)	Mean (SD)
No	298 (39.9%)	Median (Min-Max)
Unknown	322 (43.1%)	in call in the start

	(N=749)
etiology	
erulonephritis	58 (7.7%)
nder Glomerulonephritis	32 (4.3%)
ointerstitial nephritis	13 (1.7%)
ystic kidney	48 (6.4%)
tic	9 (1.2%)
etes	85 (11.3%)
ogenic Bladder	4 (0.5%)
oureteral reflux	28 (3.7%)
sis type	
toneal dialysis (PD)	50 (6.7%)
nodialysis (HD)	481 (64.2%)
PD	30 (4.0%)
emptive	188 (25.1%)
dialysis duration	
in (SD)	44.7 (57.0)
lian (Min-Max)	13.0 (0.0 - 324.0)
x CRE	
sing	117
in (SD)	7.5 (2.8%)
lian (Min-Max)	7.0 (1.2 - 20.0%)
uar Rejection	30 (4.0%)
oral Rejection	72 (9.6%)
racute Rejection	3 (0.4%)
tion	99 (13.2%)
	105 (14.0%)
loss	60 (8.0%)
w up duration (years)	
in (SD)	3.5 (2.1)
lian (Min-Max)	3.3 (0.01 - 8.0)

Gender Male 396 (52.9%) Female 353 (47.1%) Age Mean (SD) 54.1 (14.7) Median (Min-Max) 54.0 (17.0 - 89.0) BMI Missing 37 Mean (SD) 27.8 (4.9) Median (Min-Max) 27.1 (16.7 - 57.9) Blood type А 302 (40.3%) AB 37 (4.9%) В 111 (14.8%) 0 299 (39.9%) Donor type Alive 427 (57.0%) Cadeveric 322 (43.0%) HT Missing 440 Yes 117 (37.9%) No 181 (58.6%) Unknown 11 (3.6%) Diabetes Missing 440 Yes 46 (14.9%) No 253 (81.9%) Unknown 10 (3.2%)

	(N=749)
Dual kidney	13 (1.7%)
A missmatch	
0	68 (9.1%)
1	396 (52.9%)
2	285 (38.1%)
B missmatch	
0	64 (8.5%)
1	362 (48.3%)
2	323 (43.1%)
DR missmatch	
0	135 (18.0%)
1	501 (66.9%)
2	113 (15.1%)
Total missmatch	
Mean (SD)	3.6 (1.3%)
Median (Min-Max)	4 (0-6%)
Cadaver cause of death	
Missing	11
Cerebrovascular event	193 (62.1%)
Other	118 (37.9%)
Cadaveric Cold Ischemia Time (hours)	
Missing	4
Mean (SD)	19.3 (4.2%)
Median (Min-Max)	18.9 (9.2 - 39.1%)

Donor characteristics

The mean graft survival time was 6.3 95% CI [6.1-6.5].



In the multivariate joint model that incorporating both longitudinal submodels for CRE and HCT, along with a survival submodel, the risk of kidney loss increased as the age at transplantation was high (hazard ratio 1.02) and the donor type was cadaver (hazard ratio 1.08), but only the age at transplantation was found to be statistically significant (p<0.001). No effect of gender and total number of missmatches was observed (p>0.05). When the effect of longitudinal variables on survival was examined, an increase in creatinine on the logarithmic scale increased the risk, while an increase in hematocrit reduced the risk (hazard ratios 4.52 and 0.84, respectively; p<0.001).

Variable	Coefficient/SE	95% Credible Interval	P value		Coefficient/SE	95% Credible Interv
Visit				Visit	28.8476/0.5811	27.7081; 29.9835
Intercept	1.3056/0.0559	1.1960; 1.4142	<0.0001	Intercept	10.8027/0.3396	10.1493; 11.4848
$eta_{11}^{"}$	-0.6465/0.0230	-0.6916; -0.6010	<0.0001	$eta_{11}^{}$	6.7234/0.6978	5.3237; 8.0852
$\dot{\beta_{12}}$	-0.7372/0.0388	-0.8127; -0.6622	<0.0001	β_{12}	13.6235/0.6974	12.2592; 14.9774
eta_{13}	-1.4616/0.0519	-1.5612; -1.3584	<0.0001	β_{13}	5.6669/1.3665	3.0780; 8.3651
$\dot{\beta_{14}}$	0.3635/0.0791	0.2270; 0.5329	<0.0001	β_{14}	-2.2854/0.2856	-2.8480; -1.7248
Gender (ref: Male/ Female)	-0.2685/0.0267	-0.3204; -0.2164	<0.0001	Gender (ref: Male/ Female)	0.0136/0.0122	-0.0100; 0.0374
Age at Tx (years)	-0.0006/0.0011	-0.0028; 0.0017	0.6194	Age at Tx	-0.4980/0.3506	27.7081; 29.9835
Donor type (ref: Alive/Cadeveric)	0.2833/00353	0.2120; 0.3505	<0.0001	Donor type (ref: Alive/Cadeveric)	3.2892/0.0127	10.1493; 11.4848
σ_1	02921/00011	0.2900; 0.2944	<0.0001	σ_1	28.8476/0.5811	5.3237; 8.0852

Log(CRE) Longitudinal Submodel results

HCT Longitudinal Submodel results

P value

< 0.0001

<0.0001

< 0.0001

< 0.0001

< 0.0001

< 0.0001

0.2654

< 0.0001

< 0.0001

< 0.0001





Variable	Coefficient/SE	95% Credible Interval	HR	P value
CRE	1.5141 /0.2001			
нст	- 0.1696 /0.0229	-0.2149; -0.1269	0.844	<0.0001
Gender (ref: Male/ Female)	-0.0465/0.2384	-0.5264; 0.3992	0.9546	0.8646
Age at Tx (years)	0.0226 /0.0076	0.0079; 0.0379	10.229	0.0028
Donor type (ref: Alive/Cadeveric)	0.0804/0.1907	-0.2966; 0.4710	10.837	0.6752
Rejection (ref:Yok/Var)	0.0536/0.1953	-0.3384; 0.4253	10.551	0.776
Total missmatch	0.0484/0.0674	-0.0824; 0.1838	10.496	0.4744
Total dialysis duration	0.0012/0.0015	-0.0018; 0.0040	10.012	0.4212



Discussion

Studies investigating the effect of multiple longitudinal biomarkers on graft survival using joint models are limited due to computational complexity. Sayyadi et al. (Sayyadi et al., 2016) found a hazard ratio (HR) of 1.82 (95% CI 1.4-2.3; p<0.001) for graft loss associated with serum CRE levels measured monthly over one year in a survival model that included factors such as age at transplantation, donor type, immunosuppressive drug type, anemia, leukopenia, hyperkalemia, myocardial infarction, acute tubular necrosis, and anti-thymocyte globulin. Similarly, Fournier et al. (Fournier et al., 2016) found an HR of 1.89 (95% CI 1.2-3.1; p=0.01) for serum CRE measured at 3, 6, and 12 months post-transplant in a survival submodel accounting for age at transplantation, cardiovascular history, the presence of acute rejection in the first year, and donor gender.

Two studies have investigated multiple longitudinal biomarkers in joint models. The first, by Rasoul Alimi et al. (Alimi et al., 2021b)) modelled serum CRE and blood urea nitrogen with linearly in longitudinal submodels. They found an HR of 1.49 (95% CI 1.0-2.2; p=0.051) for serum CRE and 1.68 (95% CI 1.2-2.3; p<0.001) for blood urea nitrogen in a survival submodel that included age at transplantation, gender, and hypertension. Serum creatinine was not statistically significant in the presence of blood urea nitrogen. Lastly, Rizopoulos and Ghosh (Rizopoulos & Ghosh, 2011) identified eGFR, proteinuria, and hematocrit levels as being associated with graft survival in a survival submodel that included gender, age at transplantation, and weight. They modelled longitudinal biomarkes with splines in longitudinal submodels. Despite differences in risk factors and longitudinal and survival submodels, these studies consistently associated changes in serum CRE with graft survival, while hematocrit has only been associated in two studies, including the present study. Since it is known that kidney biomarkers exhibit nonlinear changes, it is important to model them using splines in joint models. This study is the second multivariate joint model to use splines for modeling kidney biomarkers and presents a realiable predictive model from a clinical perspective due to its single-center patient series.

Conclusion

The relationship between longitudinal data of kidney function biomarkers and kidney survival has been demonstrated. It is beneficial to incorporate all available information in this type of multivariate model. It is expected that dynamic predictions will be integrated into current systems in the future, with these models utilizing all data from a patient's medical history. This makes multivariate joint modeling a valuable tool in the era of personalized medicine as it provides doctors with a deeper understanding of disease progression and the ability to select the most suitable treatment at a specific follow-up time.

References

Alimi, R., Hami, M., Afzalaghaee, M., Nazemian, F., Mahmoodi, M., Yaseri, M., & Zeraati, H. (2021b). Multivariate longitudinal assessment of kidney function outcomes on graft survival after kidney transplantation using multivariate joint modeling approach: A retrospective cohort study. *Iranian Journal of Medical Sciences*, *46*(5), 364–372. https://doi.org/10.30476/ijms.2020.82857.1144

Cockwell, P., & Fisher, L.-A. (2020). The global burden of chronic kidney disease. *The Lancet, 395*(10225), 662–664. https://doi.org/10.1016/S0140-6736(19)32977-0

Fournier, M. C., Foucher, Y., Blanche, P., Buron, F., Giral, M., & Dantan, E. (2016). A joint model for longitudinal and time-to-event data to better assess the specific role of donor and recipient factors on long-term kidney transplantation outcomes. *European Journal of Epidemiology*, *31*(5), 469–479. <u>https://doi.org/10.1007/s10654-016-0121-2</u>

Rizopoulos, D., & Ghosh, P. (2011). A Bayesian semiparametric multivariate joint model for multiple longitudinal outcomes and a time-to-event. *Statistics in Medicine*, *30*(12), 1366–1380. https://doi.org/10.1002/sim.4205

Sayyadi, H., Zayeri, F., Baghestani, A. R., Baghfalaki, T., Taghizadeh Afshari, A., Mohammadrahimi, M., ... Makhdoomi, K. (2016). Assessing Risk Indicators of Allograft Survival of Renal Transplant: An Application of Joint Modeling of Longitudinal and Time-to-Event Analysis. *Iranian Red Crescent Medical Journal*, *19*(3). https://doi.org/10.5812/ircmj.40583